

Mid-~~sem~~ Examination

DA241 (2023)

Solutions and Marking Scheme.

Solution 1. Assume that the doors are labeled such that you choose Door 1.

Let us define the following events:

$S$ : You get the car following the strategy "stick to your original choice".

$C_j$ : Car is behind the door  $j$ .

By law of total probability, we condition on which door has the car:

$$P(S) = P(S|C_1) \cdot P(C_1) + \dots + P(S|C_7) \cdot P(C_7)$$

$$= \frac{1}{7} \quad \text{because:}$$

The car can be behind any door so  $P(C_j) = \frac{1}{7} \quad \forall j = 1, \dots, 7$ .

Let  $M_{i,j,k}$ : Monty opens door  $i, j$ , and  $k$ .

$$P(S) = \sum_{i,j,k} P(S|M_{i,j,k}) P(M_{i,j,k}) \quad 2 \leq i < j < k \leq 7$$

$$\text{By symmetry, } P(S|M_{i,j,k}) = P(S) = \frac{1}{7}$$

$\Rightarrow$  conditional probability that the car is behind one of the remaining 3 doors is  $\frac{6}{7}$ , which implies you should switch because then your

success probability is  $\frac{2}{7}$  rather than  $\frac{1}{7}$  if you stick with your initial

choice.

Solution 2. We are given that  $X \sim \text{Exponential}(\lambda)$ ,  $\lambda > 0$ .

$$\therefore f_X(x) = \begin{cases} \lambda e^{-\lambda x} & x \geq 0 \\ 0 & \text{o/w.} \end{cases}$$

We need to find  $E(X^3)$  using the MGF of  $X$ .

$$M(t) = E(e^{tx}) = \int_0^{\infty} e^{tx} \lambda e^{-\lambda x} dx$$

$$= \lambda \int_0^{\infty} e^{-(\lambda-t)x} dx$$

$$= \lambda \left[ \frac{e^{-(\lambda-t)x}}{-(\lambda-t)} \right]_0^{\infty}$$

$$= \lambda \left( 0 + \frac{1}{\lambda-t} \right) \quad \lambda > t$$

$$= \frac{\lambda}{\lambda-t} \quad t < \lambda$$

Now we need to use this MGF to find  $E(X^3)$ .

One way: We can take the third derivative of  $M(t)$  & evaluate it at  $t=0$ .

$$M(t) = \frac{\lambda}{\lambda-t} \quad t < \lambda$$

$$= \frac{1}{1 - \left(\frac{t}{\lambda}\right)} = \left(1 - \frac{t}{\lambda}\right)^{-1}$$

$$\frac{\partial M(t)}{\partial t} = \frac{1}{\lambda} \left(1 - \frac{t}{\lambda}\right)^{-2}$$

$$\frac{\partial^2 M(t)}{\partial t^2} = \frac{2}{1^2} \left( \frac{1-t}{1} \right)^{-3}$$

$$\frac{\partial^3 M(t)}{\partial t^3} = \frac{6}{1^3} \left( \frac{1-t}{1} \right)^{-4}$$

$$\Rightarrow E(X^3) = \left. \frac{\partial^3 M(t)}{\partial t^3} \right|_{t=0} = \frac{6}{1^3}$$

Second way: Pattern recognition:

$$M(t) = \frac{1}{1-t}$$

$$= \sum_{n=0}^{\infty} \left( \frac{t}{1} \right)^n$$

$$= \sum_{n=0}^{\infty} \frac{n! t^n}{1^n n!} \quad \text{for } t < 1$$

The  $n$ th moment of  $X$  is the coefficient of  $\frac{t^n}{n!}$  in the

$$\therefore E(X^3) = \frac{3!}{1^3} = \frac{6}{1^3}$$

Solution 3.

$X$  : # of times the code fails before the first successful run.

We recognize that  $X$  is a geometric random variable with probability  $p$  of success. — 1 mark

$$P(X=k) = q^k p \quad k=0,1,2,\dots$$

$$E(X) = \sum_{k=0}^{\infty} k q^k p.$$

$$= p \sum_{k=0}^{\infty} k q^k. \quad (*)$$

We can use the trick from the lecture:

We know, 
$$\sum_{k=0}^{\infty} q^k = \frac{1}{1-q}.$$

Taking derivative on both sides, we have

$$\sum_{k=1}^{\infty} k q^{k-1} = \frac{+1}{(1-q)^2} \quad - (**)$$

$$\Rightarrow \sum_{k=1}^{\infty} k q^k = \frac{q}{(1-q)^2}.$$

Using the above in  $(*)$ , we have:

$$E(X) = \frac{p \cdot q}{(1-q)^2} = \frac{pq}{p^2} = \frac{q}{p}. \quad - 1 \text{ mark}$$

$$\text{Var}(X) = E(X^2) - (E(X))^2$$

$$\begin{aligned} E(X^2) &= p \sum_{k=0}^{\infty} k^2 q^k \\ &= p \sum_{k=0}^{\infty} k(k-1)q^k + p \sum_{k=0}^{\infty} kq^k \end{aligned}$$

Using the same trick as before:

$$\sum_{k=2}^{\infty} k(k-1)q^{k-2} = \frac{2}{(1-q)^3} \quad \text{(taking derivative of } \frac{q}{1-q} \text{)}$$

$$\Rightarrow E(X^2) = \frac{2pq^2}{p^3} + E(X)$$

$$= \frac{2q^2}{p^2} + E(X) = \frac{2q^2}{p^2} + \frac{q}{p}$$

— 2 marks.

$$\begin{aligned} \text{Var}(X) &= \frac{2q^2}{p^2} + \frac{q}{p} - \frac{q^2}{p^2} \\ &= \frac{q^2}{p^2} + \frac{q}{p} = \frac{q^2 + qp}{p^2} = \frac{q(p+q)}{p^2} \\ &= \frac{q}{p^2} \end{aligned}$$

— 1 mark.

### Solution 1.

Let  $A_i$  :  $i$ th card in the deck has the number  $i$  written on it.

We want to find

$$P(A_1 \cup A_2 \cup \dots \cup A_n)$$
 because we win if one of

the cards has a matching number to its position.

To find the probability of the union of  $A_i$ 's, we will use inclusion-exclusion.

First,  $P(A_i) = \frac{1}{n}$  }  $P(A_i) = \frac{(n-1)!}{n!}$  - Fix  $i$ th card in  $i$ th position, rest  $n-1$  cards anywhere }

Second,  $P(A_i \cap A_j) = \frac{(n-2)!}{n!} = \frac{1}{n(n-1)}$  (Fix  $i$ th &  $j$ th positions of cards numbered  $i$  &  $j$ )

Similarly,  $P(A_i \cap A_j \cap A_k) = \frac{1}{n(n-1)(n-2)}$  following the

same pattern.

Inclusion-exclusion formula:

$$P\left(\bigcup_{i=1}^n A_i\right) = \sum_{i=1}^n P(A_i) - \sum_{i < j} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j \cap A_k) - \dots + (-1)^{n+1} P(A_1 \cap A_2 \cap \dots \cap A_n)$$

$$\therefore P\left(\bigcup_{i=1}^n A_i\right) = n \cdot \frac{1}{n} - \binom{n}{2} \frac{1}{n(n-1)} + \binom{n}{3} \frac{1}{n(n-1)(n-2)} - \dots + (-1)^{n+1} \frac{1}{n!}$$

$$= 1 - \frac{1}{2!} + \frac{1}{3!} - \dots + (-1)^{n+1} \frac{1}{n!}$$

For large  $n$ , this probability is close to  $1 - e^{-1} = 0.63$ .

$\infty$	$\infty$	} = $(\infty, \infty, \infty)$
$\infty$	$\infty$	
$\infty$	$\infty$	

Joint PDF of  $X$  and  $Y$  is given by

$$f(x, y) = \begin{cases} \frac{1}{2} & \text{if } 0 < x < 1, 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

find marginal PDF of  $X$  and  $Y$

A random point  $(x, y, z)$  is chosen uniformly in ball  $B$ .

$$B = \{ (x, y, z) : x^2 + y^2 + z^2 \leq 1 \}$$

a) The probability in three dimensions will be proportional to volume, same as it is proportional to area in two dimensions & length in one-dimension.

$$P((x, y, z) \in A) = c \cdot \text{Volume}(A) \quad A \subset B$$

where  $c$  is a constant.

$$\text{if } A = B, \quad P((x, y, z) \in B) = c \quad (\text{uniform})$$

$$\text{and } \frac{1}{c} = \frac{4\pi}{3}, \quad \text{volume of the ball.}$$

$$\therefore f(x, y, z) = \begin{cases} \frac{3}{4\pi} & \text{if } x^2 + y^2 + z^2 \leq 1 \\ 0 & \text{o/w.} \end{cases}$$

b) Joint PDF of  $X$  and  $Y$ ,

$$f_{X,Y}(x, y) = \int_{-\infty}^{\infty} f(x, y, z) dz$$

(Integrate out  $z$  from the joint PDF of  $X, Y, Z$ ).

$$\text{Since } x^2 + y^2 + z^2 \leq 1$$

$$\Rightarrow z^2 \leq 1 - x^2 - y^2$$

$$\text{classmate } \sqrt{x^2 + y^2} \leq z \leq \sqrt{1 - x^2 - y^2}$$



$$\therefore f_{X,Y}(x,y) = \int_{-\sqrt{1-x^2-y^2}}^{\sqrt{1-x^2-y^2}} \frac{3}{4\pi} dz.$$

$$= \frac{3}{4\pi} \cdot 2 \cdot \int_0^{\sqrt{1-x^2-y^2}} 1 dz$$

$$= \begin{cases} \frac{3}{2\pi} \sqrt{1-x^2-y^2} & \text{if } x^2+y^2 \leq 1 \\ 0 & \text{ofw.} \end{cases}$$

c). From part b), we can integrate out  $y$  from  $f_{X,Y}(x,y)$  to get marginal of  $X$ .

$$f_X(x) = \frac{3}{2\pi} \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \sqrt{1-x^2-y^2} dy.$$

$$-1 \leq x \leq 1.$$